

Are crude Oil, Gas and Coal Prices Cointegrated?

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Abstract

This study examines the existence of long run relation between crude oil, natural gas and coal prices. Energy data for US is used and Based on the result of The Augmented Dickey-Fuller (ADF) tests, autoregressive distributed lag (ARDL) approach is adapted to cointegration analysis. Underlying ARDL model is specified in logarithmic form, so that the coefficients indicate the elasticities. Long run relationship and error correction model (ECM) are estimated for selected ARDL. Moreover, to confirm the stability of the model, CUSUM and CUSUMSQ tests are also conducted with the results that the estimated model is completely stable. The results confirm the existence of long run relation between coal, gas and oil prices. However, in short run gas prices have no effects on the oil prices as its coefficient is insignificant.

Keywords: energy (Oil, Gas, Coal) prices; unit root; cointegration; autoregressive distributed lag (ARDL; ECM model).

1- Introduction

One characteristic of commodity prices is the presence of a unit root in their univariate time series representation, implying that price movements are better characterized as being the sum of permanent and transitory components where the permanent component is a random walk.

However, this study is to examine the long-run relation between crude oil, gas and coal prices. In doing so, tests for unit roots in the univariate time series representation of monthly prices are performed to determine their integration degrees- a prerequisite for the analysis of cointegration. Cointegration is designed to deal explicitly with the analysis of the

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relationship between non-stationary time series. In particular, it allows individual time series to be non-stationary but requires a linear combination of the series to be stationary. Therefore, the basic idea behind cointegration is to search for linear combinations of individually non-stationary time series that are themselves stationary. The methodology used to study common trends in these series is based on the autoregressive distributed lag (ARDL) approach proposed by Pesaran and Shin (Pesaran and Shin, (1995, 1998); Pesaran et al., (1996); Pesaran et al. (2001)).

This paper is developed in six sections. Section two is a short literature review. Section three reports the results of unit root and stationarity tests using Augmented Dickey-Fuller (ADF) approach. Section four describes the methodology of autoregressive distributed lag (ARDL) approach. Section five presents empirical results and the last section reports the main conclusions.

2- Literature Review

What price linkages exist between oil, coal, and natural gas? The main theory here is based on economic substitution. As Bachmeier & Griffin (2006) assert clearly, in the short run coal, oil, and natural gas are not fungible and direct fuel competition is limited. In the long run, these fuels become much closer economic substitutes depending on their respective costs of conversion technologies (Griffin (1979)).

It likes to be probable conjecture that long run relationship between oil, gas and coal price exist. But existence of short run relationship is doubtful.

Serletis (1992) tested for unit roots in the univariate time-series representation of the daily crude oil, heating oil, and unleaded gasoline spot-month future prices. The results showed that the random walk hypothesis for daily energy future prices can be rejected if allowance is made for the possibility of a one-time break in the intercept and the slope of the trend function at an unknown point in time.

Serletis (1994) reported that the maximum likelihood cointegration analysis of daily spot-month crude oil, heating oil and unleaded gasoline future prices covering the period 3 December 1984 to 30 April 1993 led to the conclusion that all three spot-month future prices are driven by only one common trend, suggesting that it is appropriate to model energy future prices as a cointegrated system.

De Vany and Walls (1995) analyzed the degree of integration of the North American gas market and the way price dynamics evolved as these markets were progressively embedded in a larger web of open pipelines and interconnected markets. With the two-step Engle-Granger test for cointegration, spot prices were found to be increasingly co-integrated as open access to the pipelines expands through the network.

King and Cuc (1996) investigated the strength of spot price integration between various natural gas producing basins of North America, from the mid 1980's until the mid 1990's and with time varying parameter (Kalman Filter) and cointegration analysis. Bivariate cointegration tests (Engle-Granger procedure) results were qualitatively similar to De Vany and Walls (1995). Time varying parameter analysis results indicated that price convergence has been emerging in regional markets.

Serletis (1997), in a slightly different manner, tested for shared stochastic trends in the North American markets. Evidence concerning the shared stochastic trends in eight North American natural gas spot markets, using monthly data (1990: 06 - 1996: 01), was obtained by the Engle-Granger approach and the Johansen maximum likelihood approach. Prices within eastern and western areas were found to be driven by different stochastic trends.

Serletis and Herbert (1999) investigated the dynamics of North American natural gas, fuel oil and power prices in the area of eastern Pennsylvania, New Jersey, Maryland and Delaware, using daily data (1996: 10 - 1997: 11) on the Henry Hub and Transco Zone 6 natural gas prices, the PJM (Pennsylvania, New-Jersey and Maryland) power market for electricity price and the fuel oil price for New York Harbor. Correlation between prices in log levels was first investigated and the stationary properties of the prices were analysed using the ADF test. The Engle-Granger Bivariate cointegration test for the pairs of integrated series reported that each pair cointegrates, leading to the conclusion that the same underlying stochastic component affects the three markets.

Asche et al. (2000) investigated the degree of market integration for France, Germany and Belgium. Cointegration tests highlighted that the different border prices for gas to France move proportionally over time and without any significant differences in mean. Furthermore, national markets in Germany, France and Belgium were found to be highly integrated.

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Asche et al. (2002) examined whether the German market was integrated by investigating time series of Norwegian, Dutch and Russian Gas monthly export prices to Germany from January 1990 to December 1997. The Johansen multivariate procedure results showed that gas from the three suppliers compete closely in the same markets since the prices move proportionally over time, but at different price levels.

Hirschhausen et al. (2004) examined the degree of natural gas market integration in Europe, North America and Japan, between the mid 1990's and 2002. Corresponding hypothesis was that there was a certain split of prices between Europe and North America. The relationship between the international gas market prices and their relation to the oil price, are investigated through principal component analysis and Johansen likelihood-based procedures. Both of them showed a high level of integration within the European/Japanese and North American markets and that the European/Japanese and the North American markets are connected to a much lesser extent.

Warell (2006) tested the hypothesis of the existence of a single economic market for the international coal industry, separated for coking and steam coal, and has investigated market integration over time. This were conducted by applying cointegration and error-correction models on quarterly price series data in Europe and Japan over the time period 1980-2000. Both the coking and the steam coal markets showed evidence of global market integration, as demonstrated by the stable long-run cointegrating relationship between the respective price series in different world regions.

Bachmeier and Griffin (2006) evaluated the degree of market integration both within and between crude oil, coal, and natural gas markets. Their approach yields parameters that can be readily tested against a priori conjectures. Using daily price data for five very different crude oils, they concluded that the world oil market is a single, highly integrated economic market. On the other hand, coal prices at five trading locations across the United States are cointegrated, but the degree of market integration is much weaker, particularly between Western and Eastern coals. Finally, they showed that crude oil, coal, and natural gas markets are only very weakly integrated.

Theodore and Emilie (2007) examined the relationship between UK wholesale gas prices and the Brent oil price over the period 1996–2003 in

order to investigate whether oil and gas prices ‘decoupled’ during this period as orthodox gas market liberalisation theory had suggested. Tests for unit roots and cointegration were carried out and it was discovered that a long-run equilibrium relationship between UK gas and oil prices exists. It was found that the cointegrating relationship is present throughout the sample period. However, the long-run solutions seem to be more volatile. Evidence was provided that the short-run relationship is linear and impulse response functions are used to examine the effects that a shock in oil would have on gas. These findings do not support the assumption that gas prices and oil prices ‘decouple’.

Hammoudeh et al. (2008) examined the dynamic relationship between pairs of four oil benchmark prices (i.e., West Texas Intermediate, Brent, Dubai, and Maya). The results indicated that there is a long-run equilibrium relationship between different benchmarks, regardless of their properties and locations.

3- The Data and Unit Roots

We used data for US energy markets. The Crude Oil Domestic First Purchase Price, the natural gas Wellhead Price and the Cost of Coal Receipts at Electric Generating Plants for coal price were employed. The time period of the analysis extends from October 1983 to October 2008, involving 301 observations¹.

The results of unit roots tests reported by Microfit based on Schwarz Bayesian Criterion (SBC) are as follow:

Table 1: Augmented Dickey-Fuller (ADF) test

variables	an intercept but not a trend		an intercept and a linear trend	
	Critical values	Statistic	Critical values	Statistic
Oil	-2.8714	-2.8683	-3.4264	-3.7656
Gas	-2.8714	-1.5382	-3.4264	-3.6957
Coal	-2.8714	3.0999	-3.4264	3.2806
D1Oil	-2.8718	-5.2030	-3.4270	-5.6448
D1Gas	-2.8718	-5.3603	-3.4270	-5.6208

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D1Coal	-2.8718	1.1232	-3.4270	-.72409
D2Coal	-2.8722	-1.2675	-3.4276	-1.5834
D3Coal	-2.8727	-1.4362	-3.4283	-1.4883
D4Coal	-2.8732	-3.1544	-3.4291	-3.3907
D5Coal	-2.8738	-7.9277	-3.4300	-8.0053

Based on the results of ADF test, oil, gas and coal are integrated processes of degree one, one and five or I (1), I (1) and I (5), respectively.

4- Methodology

There are various techniques for conducting the cointegration analysis. Econometric literature has abundant econometric techniques to investigate cointegration relationships among economic variables. The popular approaches are the well-known residual based approach proposed by Engle and Granger (1987) and the maximum likelihood-based approach proposed by Johansen and Julius (1990) and Johansen (1992).

In applying the cointegration technique, we need to determine the order of integration for each variable. When there are more than two I (1) variables in the system, the maximum likelihood approach of Johansen and Julius has the advantage over residual-based approach of Engle and Granger; however, both of the approaches require that the variables have the same order of integration. This requirement often causes difficulty to the researchers when the system contains the variables with different orders of integration- such as in this study. To overcome this problem, Pesaran et al. (1996, 2001) proposed a new approach known as Autoregressive Distributed Lag (ARDL) for cointegration test that does not require the classification of variables into I(0) or I(1). More recent studies have indicated that the ARDL approach to cointegration is preferable to other conventional cointegration approaches. The ARDL is applicable irrespective of whether the underlying regressors are purely I(0), purely I(1) or mutually cointegrated. The statistic underlying the procedure is the Wald or F-statistic in a generalized Dickey–Fuller type regression, which is used to test the significance of lagged levels of the variables under consideration in a conditional unrestricted equilibrium correction model (ECM) (Pesaran et al, 2001). Besides, ARDL approach is more robust and performs better for small sample sizes than other cointegration techniques.

The ARDL approach involves estimating the conditional error correction version of the ARDL model. The augmented ARDL (p, q₁, q₂, ..., q_k) is given by the following equation (Pesaran and Pesaran, 1997; Pesaran and Shin, 2001):

$$\alpha(L, p)y_t = \alpha_0 + \sum_{i=1}^k \beta_i(L, q_i)x_{it} + \lambda w_t + \varepsilon_t \quad \forall i=1, 2, \dots, n \quad (3)$$

Where

$$\alpha(L, p) = 1 - \alpha_1 L - \alpha_2 L^2 - \dots - \alpha_p L^p$$

$$\beta_i(L, q_i) = \beta_{i0} - \beta_{i1} L - \beta_{i2} L^2 - \dots - \beta_{iq_i} L^{q_i} \quad \forall i = 1, 2, \dots, k$$

y_t is the dependent variable, α_0 is the constant term, L is the lag operator, w_t is $1 \times s$ vector of deterministic variables such as intercept term, time trends, or exogenous variables with fixed lags. x_t is the k -dimensional forcing variables which are not cointegrated among themselves. ε_t is a vector of stochastic error terms, with zero means and constant variance-covariance.

The long-run coefficients are estimated by:

$$\pi = \frac{\hat{\lambda}(\hat{p}, \hat{q}_1, \hat{q}_2, \dots, \hat{q}_k)}{1 - \hat{\alpha}_1 - \hat{\alpha}_2 - \dots - \hat{\alpha}_{\hat{p}}}$$

where $\hat{\lambda}(\hat{p}, \hat{q}_1, \hat{q}_2, \dots, \hat{q}_k)$ denotes the OLS estimates of λ' in equation (1) for the selected ARDL model.

The error correction model (ECM) related to the ARDL($\hat{p}, \hat{q}_1, \hat{q}_2, \dots, \hat{q}_k$) can be obtained by writing equation (3) in terms of lagged levels and the first difference of $y_t, x_{1t}, x_{2t}, \dots, x_{kt}$ and w_t :

$$\Delta y_t = \Delta \alpha_0 - \alpha(1, \hat{p}) ECM_{t-1} + \sum_{i=1}^k \beta_{i0} x_{it} + \lambda \Delta w_t - \sum_{j=1}^{\hat{p}-1} \alpha_j \Delta y_{t-j} - \sum_{i=1}^k \sum_{j=1}^{\hat{q}_i-1} \beta_{ij} \Delta x_{i,t-j} + \varepsilon_t \quad (5)$$

where ECM is the error correction model and it is defined as follows:

$$ECM_t = y_t - \hat{\alpha} - \sum \hat{\beta}_{i0} x_{it} - \lambda' w_t$$

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The ARDL approach involves two steps for estimating the long-run relationship (Pesaran et al., 2001). The first step is to examine the existence of long-run relationship among all variables in the equations under estimation. The second step is to estimate the long-run and the short-run coefficients of the same equation. We run the second step only if we find a long-run relationship in the first step (Narayan, et al. 2004).

5- Autoregressive Distributed Lag (ARDL) Model and Empirical Results

Because underlying variables (Oil, Gas and Coal) have different integration degrees, ARDL approach is adopted for cointegration analysis.

Our ARDL model to estimate long-run relationship among underlying variables, in logarithmic form, is specified as follows:

$$LOil_t = \alpha_0 + \alpha_1 t + \sum_{i=1}^{p-1} \beta_i LOil_{t-i} + \sum_{j=0}^{q-1} \theta_{1j} LGas_{t-j} + \sum_{j=0}^{q-1} \theta_{2j} LCoal_{t-j} + \varepsilon_t \quad (7)$$

Where α_0 is the constant term, and t is time trend.

Using Microfit for estimation, ARDL (2, 0, 0) is selected based on Schwarz Bayesian Criterion (SBC). Table (2) shows the results for the selected ARDL model.

**Table 2: Autoregressive Distributed Lag Estimates
ARDL (2,0,0) selected based on Schwarz Bayesian Criterion**

Regressor	Coefficient	T-Ratio	Standard Error
LOil(-1)	1.3854	26.2829	.052712
LOil(-2)	-.48740	-9.1044	.053534
LGas	.039038	1.7886	.021826
LCoal	.021046	3.4231	.061481
Constant	.15225	3.5821	2.3791
Time trend	.32883	3.1119	.10563
R-Squared = .98078		R-Bar-Squared = .98044	
F-statistic F(5, 281) = 2868		Schwarz Bayesian Criterion = 333.9134	

As the table indicates all the statistics of the estimated model, except for gas price coefficient, are satisfactory.

To confirm the stability of the estimated model, the tests of Cumulative Sum of Recursive Residuals (CUSUM) and Cumulative Sum of Squares of Recursive Residuals (CUSUMSQ) are employed in this paper. Figures (1) and (2) provide the graphs of CUSUM and CUSUMSQ tests, respectively.

These figures indicate that the plots of CUSUM and CUSUMQ are completely stable within 5% of critical bands.

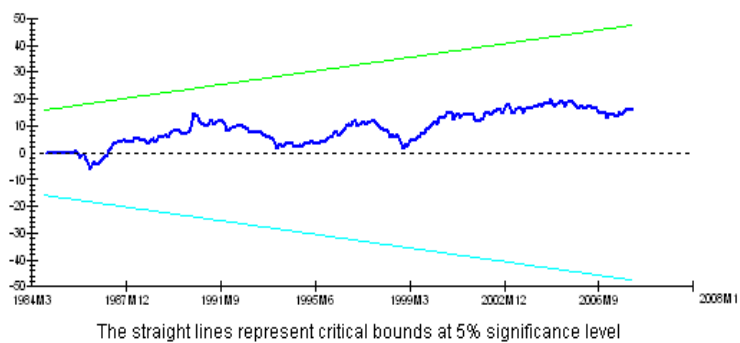


Figure 1: Cumulative Sum of Recursive Residuals (CUSUM)

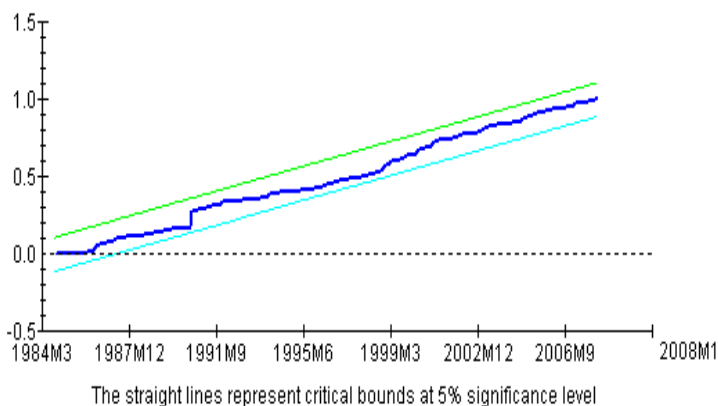


Figure 2: Cumulative Sum of Squares of Recursive Residuals (CUSUMSQ)

The results of estimated long run regression and the corresponding ECM model are reported in tables (3) and (4).

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**Table 3: Estimated Long Run Coefficients using the ARDL Approach
ARDL (2, 0, 0) selected based on Schwarz Bayesian Criterion**

Regressor	Coefficient	T-Ratio	Standard Error
LGas	.38281	2.2419	.17076
LCoal	2.0638	4.2164	.48946
Constant	1.4930	7.8976	.18904
Time trend	.0032238	2.9587	.0010896

**Table (4): Error Correction Representation for the Selected ARDL Model
ARDL (2,0,0) selected based on Schwarz Bayesian Criterion**

Regressor	Coefficient	T-Ratio	Standard Error
D2LOil	.48740	9.1044	.053534
D1LGas	.039038	1.7886	.021826
D1LCoal	.21046	3.4231	.061481
DConstant	.15225	3.5821	.042502
DTime trend	.32883	3.1119	.10563
ECM(-1)	-.10198	-4.4794	.022765
R-Squared = .25578		R-Bar-Squared = .24254	
F-statistic F(5, 281) 19.3156		Schwarz Bayesian Criterion = 333.9134	

As table (3) indicates, the coefficients (Gas prices and Coal prices) are statistically significant.

From table (4), it is clear that the error correction term (ECM (-1)) has the right sign (negative) and is statistically significant. Specifically, the estimated value of EC_{t-1} is -.10198.

The absolute value of the coefficient of EC_{t-1} , indicating the speed of adjustment to equilibrium, denotes that 10% of any shocks dumps out in each period, converging back to the long run equilibrium. Bannerjee, Dolado and Mestre (1998), hold that a highly significant error correction term is further proof of the existence of a stable long-term relationship, which is the case here.

Except coefficient of D1LGas (first difference of LGas), all other short run coefficients are significant with positive signs as it is expected.

6- Conclusion

This paper investigates the existence of long-run relation between crude oil, gas and coal prices. The data for us energy market is used in this study. The time period of the analysis extends from October 1983 to October 2008, involving 301 observations. Augmented Dickey-Fuller (ADF) approach

employed for existence of unit root. In brief, Augmented Dickey-Fuller (ADF) tests indicate that Oil and Gas prices are integrated processes of degree one or I (1) and Coal prices is integrated processes of degree five or I (5). As integration degree of variables are not same the Autoregressive Distributed Lag (ARDL) approach to cointegration adapted to cointegration analysis on Oil, Gas and coal prices. The ARDL model was specified in logarithmic form which coefficients mean as elasticities. The model selection fulfilled by Schwarz Bayesian Criterion (SBC) and so ARDL (2, 0, 0) was selected. Moreover, to confirm the stability of the model, CUSUM and CUSUMSQ tests are also conducted with the results that the estimated model is completely stable.

At short run coefficient of D1LGas (first difference of LGas), are statistically insignificant but in long run LGas (logarithm of gas prices) have a significant coefficient. LCoal (logarithm of coal prices) at long run and D1LCoal (first difference of LCoal) at short run have significant coefficients. The results provide that second difference of LOil (D2LOil), indicating short run own price elasticity of oil prices, have a significant coefficient.

The estimated value of EC_{t-1} is -.10198 indicating about 10% speed of adjustment toward equilibrium. It is clear from estimated value of EC_{t-1} that the error correction term (ECM (-1)) has the right sign (negative) and is statistically significant.

Based on the result of cointegration analysis, we can found the long run relationship between oil, gas and coal prices.

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Appendix.1

Unit root tests for variable OIL
The Dickey-Fuller regressions include an intercept but not a trend

297 observations used in the estimation of all ADF regressions.
Sample period from 1984M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-1.24703	-749.8791	-751.8791	-755.5728	-753.3578
ADF(1)	-2.8683	-694.2077	-697.2077	-702.7483	-699.4258
ADF(2)	-2.7050	-694.2037	-698.2037	-705.5912	-701.1612
ADF(3)	-1.8989	-693.1477	-698.1477	-707.3820	-701.8445

95% critical value for the augmented Dickey-Fuller statistic = -2.8714
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable OIL
The Dickey-Fuller regressions include an intercept and a linear trend

297 observations used in the estimation of all ADF regressions.
Sample period from 1984M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-1.6015	-747.4451	-750.4451	-755.9857	-752.6632
ADF(1)	-3.7656	-691.2862	-695.2862	-702.6736	-698.2436
ADF(2)	-3.6319	-691.2852	-696.2852	-705.5195	-699.9820
ADF(3)	-2.9315	-690.3775	-696.3775	-707.4587	-700.8137

95% critical value for the augmented Dickey-Fuller statistic = -3.4264
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable GAS
The Dickey-Fuller regressions include an intercept but not a trend

297 observations used in the estimation of all ADF regressions.
Sample period from 1984M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-1.5854	-182.9664	-184.9664	-188.6602	-186.4452
ADF(1)	-1.9937	-178.7620	-181.7620	-187.3026	-183.9801
ADF(2)	-1.5382	-175.3503	-179.3503	-186.7378	-182.3078
ADF(3)	-1.8087	-174.2769	-179.2769	-188.5112	-182.9737

95% critical value for the augmented Dickey-Fuller statistic = -2.8714
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable GAS
The Dickey-Fuller regressions include an intercept and a linear trend

297 observations used in the estimation of all ADF regressions.
Sample period from 1984M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-3.1691	-179.0640	-182.0640	-187.6046	-184.2821
ADF(1)	-3.6957	-173.8772	-177.8772	-185.2647	-180.8347
ADF(2)	-3.1914	-171.2459	-176.2459	-185.4802	-179.9427
ADF(3)	-3.4902	-169.6975	-175.6975	-186.7787	-180.1337

95% critical value for the augmented Dickey-Fuller statistic = -3.4264
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable COAL

The Dickey-Fuller regressions include an intercept but not a trend

297 observations used in the estimation of all ADF regressions.
Sample period from 1984M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	3.0999	748.4810	746.4810	742.7873	745.0023
ADF(1)	2.8249	748.6203	745.6203	740.0797	743.4022
ADF(2)	2.8203	748.6997	744.6997	737.3122	741.7422
ADF(3)	2.6287	748.8187	743.8187	734.5844	740.1219

95% critical value for the augmented Dickey-Fuller statistic = -2.8714
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable COAL

The Dickey-Fuller regressions include an intercept and a linear trend

297 observations used in the estimation of all ADF regressions.
Sample period from 1984M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	3.2806	760.3759	757.3759	751.8353	755.1578
ADF(1)	3.4057	760.8059	756.8059	749.4184	753.8484
ADF(2)	3.9004	762.9563	757.9563	748.7219	754.2594
ADF(3)	4.0978	763.7465	757.7465	746.6653	753.3103

95% critical value for the augmented Dickey-Fuller statistic = -3.4264
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D1OIL

The Dickey-Fuller regressions include an intercept but not a trend

285 observations used in the estimation of all ADF regressions.
Sample period from 1985M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-2.9198	-785.4580	-787.4580	-791.1105	-788.9222
ADF(1)	-5.2030	-738.7413	-741.7413	-747.2200	-743.9376
ADF(2)	-5.0599	-738.6684	-742.6684	-749.9733	-745.5967
ADF(3)	-5.4320	-736.8050	-741.8050	-750.9362	-745.4654

95% critical value for the augmented Dickey-Fuller statistic = -2.8718
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D1OIL

The Dickey-Fuller regressions include an intercept and a linear trend

285 observations used in the estimation of all ADF regressions.
Sample period from 1985M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-3.1738	-784.6159	-787.6159	-793.0946	-789.8122
ADF(1)	-5.6448	-736.2291	-740.2291	-747.5341	-743.1575
ADF(2)	-5.5467	-736.0230	-741.0230	-750.1542	-744.6835
ADF(3)	-5.9900	-733.6162	-739.6162	-750.5736	-744.0087

95% critical value for the augmented Dickey-Fuller statistic = -3.4270
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

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Unit root tests for variable D1COAL
The Dickey-Fuller regressions include an intercept but not a trend

285 observations used in the estimation of all ADF regressions.
Sample period from 1985M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	1.1232	739.2722	737.2722	733.6197	735.8080
ADF(1)	1.7434	741.8348	738.8348	733.3561	736.6385
ADF(2)	1.9177	742.1895	738.1895	730.8845	735.2611
ADF(3)	1.4291	743.4307	738.4307	729.2995	734.7702

95% critical value for the augmented Dickey-Fuller statistic = -2.8718
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D1COAL
The Dickey-Fuller regressions include an intercept and a linear trend

285 observations used in the estimation of all ADF regressions.
Sample period from 1985M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-.72409	741.4587	738.4587	732.9800	736.2624
ADF(1)	-.11727	743.5638	739.5638	732.2588	736.6354
ADF(2)	.079055	743.7695	738.7695	729.6382	735.1090
ADF(3)	-.34118	745.2621	739.2621	728.3047	734.8695

95% critical value for the augmented Dickey-Fuller statistic = -3.4270
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D3COAL

The Dickey-Fuller regressions include an intercept but not a trend

 261 observations used in the estimation of all ADF regressions.
 Sample period from 1987M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-4.0029	425.6734	423.6734	420.1089	422.2406
ADF(1)	-2.6264	430.7810	427.7810	422.4343	425.6318
ADF(2)	-1.4362	435.7809	431.7809	424.6519	428.9153
ADF(3)	-1.3700	435.7810	430.7810	421.8697	427.1990

 95% critical value for the augmented Dickey-Fuller statistic = -2.8727
 LL = Maximized log-likelihood AIC = Akaike Information Criterion
 SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D3COAL

The Dickey-Fuller regressions include an intercept and a linear trend

 261 observations used in the estimation of all ADF regressions.
 Sample period from 1987M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-4.0726	426.2911	423.2911	417.9443	421.1418
ADF(1)	-2.6925	431.5275	427.5275	420.3985	424.6619
ADF(2)	-1.4883	436.7620	431.7620	422.8507	428.1799
ADF(3)	-1.3893	436.7659	430.7659	420.0724	426.4675

 95% critical value for the augmented Dickey-Fuller statistic = -3.4283
 LL = Maximized log-likelihood AIC = Akaike Information Criterion
 SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

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Unit root tests for variable D2COAL
The Dickey-Fuller regressions include an intercept but not a trend

273 observations used in the estimation of all ADF regressions.
Sample period from 1986M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-2.3969	589.1884	587.1884	583.5790	585.7395
ADF(1)	-1.2675	593.2946	590.2946	584.8804	588.1213
ADF(2)	-.52405	595.1762	591.1762	583.9573	588.2784
ADF(3)	-.88412	596.0268	591.0268	582.0031	587.4045

95% critical value for the augmented Dickey-Fuller statistic = -2.8722
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D2COAL
The Dickey-Fuller regressions include an intercept and a linear trend

273 observations used in the estimation of all ADF regressions.
Sample period from 1986M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-2.7048	590.3310	587.3310	581.9168	585.1576
ADF(1)	-1.5834	594.3125	590.3125	583.0935	587.4147
ADF(2)	-.84381	596.1512	591.1512	582.1275	587.5289
ADF(3)	-1.1951	597.0454	591.0454	580.2170	586.6987

95% critical value for the augmented Dickey-Fuller statistic = -3.4276
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D4COAL
 The Dickey-Fuller regressions include an intercept but not a trend

 249 observations used in the estimation of all ADF regressions.
 Sample period from 1988M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-5.9504	265.3100	263.3100	259.7925	261.8942
ADF(1)	-4.4875	269.0223	266.0223	260.7461	263.8986
ADF(2)	-3.1544	274.6810	270.6810	263.6461	267.8493
ADF(3)	-2.8805	274.7770	269.7770	260.9834	266.2375

 95% critical value for the augmented Dickey-Fuller statistic = -2.8732
 LL = Maximized log-likelihood AIC = Akaike Information Criterion
 SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D4COAL
 The Dickey-Fuller regressions include an intercept and a linear trend

 249 observations used in the estimation of all ADF regressions.
 Sample period from 1988M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-6.1771	266.7111	263.7111	258.4349	261.5873
ADF(1)	-4.7126	270.2383	266.2383	259.2034	263.4066
ADF(2)	-3.3907	275.8762	270.8762	262.0826	267.3367
ADF(3)	-3.1113	275.9763	269.9763	259.4239	265.7288

 95% critical value for the augmented Dickey-Fuller statistic = -3.4291
 LL = Maximized log-likelihood AIC = Akaike Information Criterion
 SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D1GAS
 The Dickey-Fuller regressions include an intercept but not a trend

 285 observations used in the estimation of all ADF regressions.
 Sample period from 1985M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-4.8524	-270.8154	-272.8154	-276.4679	-274.2796
ADF(1)	-5.4135	-267.5865	-270.5865	-276.0653	-272.7828
ADF(2)	-4.5124	-265.2954	-269.2954	-276.6004	-272.2238
ADF(3)	-5.3603	-259.6785	-264.6785	-273.8097	-268.3389

 95% critical value for the augmented Dickey-Fuller statistic = -2.8718
 LL = Maximized log-likelihood AIC = Akaike Information Criterion
 SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable D1GAS
 The Dickey-Fuller regressions include an intercept and a linear trend

 285 observations used in the estimation of all ADF regressions.
 Sample period from 1985M2 to 2008M10

	Test Statistic	LL	AIC	SBC	HQC
DF	-5.0562	-269.8363	-272.8363	-278.3150	-275.0325
ADF(1)	-5.6607	-266.2848	-270.2848	-277.5898	-273.2132
ADF(2)	-4.7465	-264.2367	-269.2367	-278.3680	-272.8972
ADF(3)	-5.6208	-258.3087	-264.3087	-275.2662	-268.7013

 95% critical value for the augmented Dickey-Fuller statistic = -3.4270
 LL = Maximized log-likelihood AIC = Akaike Information Criterion
 SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

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Appendix.2

```
Unit root tests for variable D5COAL
The Dickey-Fuller regressions include an intercept but not a trend
*****
237 observations used in the estimation of all ADF regressions.
Sample period from 1989M2 to 2008M10
*****
      Test Statistic      LL      AIC      SBC      HQC
DF      -7.9277      114.4020      112.4020      108.9340      111.0042
ADF(1)  -6.3136      115.8504      112.8504      107.6483      110.7537
ADF(2)  -4.8513      119.7881      115.7881      108.8519      112.9924
ADF(3)  -4.5798      119.7900      114.7900      106.1199      111.2954
*****
95% critical value for the augmented Dickey-Fuller statistic = -2.8738
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion      HQC = Hannan-Quinn Criterion
```

```
Unit root tests for variable D5COAL
The Dickey-Fuller regressions include an intercept and a linear trend
*****
237 observations used in the estimation of all ADF regressions.
Sample period from 1989M2 to 2008M10
*****
      Test Statistic      LL      AIC      SBC      HQC
DF      -8.0053      115.0015      112.0015      106.7994      109.9047
ADF(1)  -6.3848      116.3175      112.3175      105.3814      109.5218
ADF(2)  -4.9127      120.1129      115.1129      106.4428      111.6183
ADF(3)  -4.6446      120.1133      114.1133      103.7091      109.9198
*****
95% critical value for the augmented Dickey-Fuller statistic = -3.4300
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion      HQC = Hannan-Quinn Criterion
```

```
Estimated Long Run Coefficients using the ARDL Approach
ARDL(2,0,0) selected based on Schwarz Bayesian Criterion
*****
Dependent variable is LOIL
287 observations used for estimation from 1984M3 to 2008M1
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
LGAS      .38281      .17076      2.2419[.026]
LCOAL      2.0638      .48946      4.2164[.000]
A      1.4930      .18904      7.8976[.000]
T      .0032238      .0010896      2.9587[.003]
*****
```

Unit root tests for variable RESIDUAL
The Dickey-Fuller regressions include an intercept but not a trend

283 observations used in the estimation of all ADF regressions.
Sample period from 1984M7 to 2008M1

	Test Statistic	LL	AIC	SBC	HQC
DF	-15.8592	344.8493	342.8493	339.2039	341.3876
ADF(1)	-12.8338	346.5386	343.5386	338.0705	341.3461
ADF(2)	-9.6537	346.9300	342.9300	335.6391	340.0066
ADF(3)	-8.7120	347.1841	342.1841	333.0705	338.5299

95% critical value for the augmented Dickey-Fuller statistic = -2.8719
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

Unit root tests for variable RESIDUAL
The Dickey-Fuller regressions include an intercept and a linear trend

283 observations used in the estimation of all ADF regressions.
Sample period from 1984M7 to 2008M1

	Test Statistic	LL	AIC	SBC	HQC
DF	-15.8326	344.8653	341.8653	336.3971	339.6728
ADF(1)	-12.8131	346.5577	342.5577	335.2668	339.6343
ADF(2)	-9.6384	346.9489	341.9489	332.8353	338.2946
ADF(3)	-8.6978	347.2024	341.2024	330.2660	336.8173

95% critical value for the augmented Dickey-Fuller statistic = -3.4271
LL = Maximized log-likelihood AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion HQC = Hannan-Quinn Criterion

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```

Autoregressive Distributed Lag Estimates
ARDL(2,0,0) selected based on Schwarz Bayesian Criterion
*****
Dependent variable is LOIL
287 observations used for estimation from 1984M3 to 2008M1
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
LOIL(-1)           1.3854           .052712             26.2829[.000]
LOIL(-2)           -.48740          .053534             -9.1044[.000]
LGAS                .039038          .021826             1.7886[.075]
LCOAL              .21046           .061481             3.4231[.001]
A                  .15225           .042502             3.5821[.000]
T                  .3288E-3         .1056E-3            3.1119[.002]
*****
R-Squared          .98078           R-Bar-Squared       .98044
S.E. of Regression .072007         F-stat.             F( 5, 281) 2868.0[.000]
Mean of Dependent Variable 3.0344         S.D. of Dependent Variable .51485
Residual Sum of Squares 1.4570         Equation Log-likelihood 350.8919
Akaike Info. Criterion 344.8919       Schwarz Bayesian Criterion 333.9134
DW-statistic       1.8840
*****

```

```

Diagnostic Tests
*****
* Test Statistics *      LM Version      *      F Version      *
*****
* A:Serial Correlation*CHSQ( 12)= 12.4026[.414]*F( 12, 269)= 1.0125[.437]*
*
* B:Functional Form *CHSQ( 1)= 2.5460[.111]*F( 1, 280)= 2.5061[.115]*
*
* C:Normality *CHSQ( 2)= 111.4867[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= .66025[.416]*F( 1, 285)= .65716[.418]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

Error Correction Representation for the Selected ARDL Model
 ARDL(2,0,0) selected based on Schwarz Bayesian Criterion

Dependent variable is dLOIL
 287 observations used for estimation from 1984M3 to 2008M1

Regressor	Coefficient	Standard Error	T-Ratio[Prob]
dLOIL1	.48740	.053534	9.1044[.000]
dLGAS	.039038	.021826	1.7886[.075]
dLCOAL	.21046	.061481	3.4231[.001]
dA	.15225	.042502	3.5821[.000]
dT	.3288E-3	.1056E-3	3.1119[.002]
ecm(-1)	-.10198	.022765	-4.4794[.000]

List of additional temporary variables created:

dLOIL = LOIL-LOIL(-1)
 dLOIL1 = LOIL(-1)-LOIL(-2)
 dLGAS = LGAS-LGAS(-1)
 dLCOAL = LCOAL-LCOAL(-1)
 dA = A-A(-1)
 dT = T-T(-1)
 ecm = LOIL -.38281*LGAS -2.0638*LCOAL -1.4930*A -.0032238*T

R-Squared	.25578	R-Bar-Squared	.24254
S.E. of Regression	.072007	F-stat.	F(5, 281) 19.3156[.000]
Mean of Dependent Variable	.0042028	S.D. of Dependent Variable	.082736
Residual Sum of Squares	1.4570	Equation Log-likelihood	350.8919
Akaike Info. Criterion	344.8919	Schwarz Bayesian Criterion	333.9134
DW-statistic	1.8840		

R-Squared and R-Bar-Squared measures refer to the dependent variable
 dLOIL and in cases where the error correction model is highly
 restricted, these measures could become negative.