Crowding out Effect of Foreign Aid in Selected Developing Countries: Panel Data Evidence

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Abstract

Public sector decision-makers are faced with the task of allocating resources among different alternative subject due budgetary constraints. In this paper Official Development Assistance (ODA) data have been considered as foreign aid. ODA is channeled through the public sector of recipient countries and, hence, the ultimate effect of ODA on savings or economic growth depends on how governments respond to it. This paper tries to explore the impact of official development assistance on public sector behavior in selected developing countries and contribute to the fiscal response literature on two main grounds. First, it specifies a fiscal response model. Second, using panel data model for a sample of 25 aid recipient countries in Asia and Latin America over the period 1991-2010. Empirical results indicate that official development assistance has a positive and significant effect on government investment expenditure, but it has not significantly impact on government current expenditures. Results also show that ODA crowds out both government revenue and public borrowing.

Keywords: Public Fiscal Behavior, Official Development Assistance, Panel Data.

1- Introduction

Based on the department of Economic and Social Affairs of UN for many low income countries, Official Development Assistance remains an important vehicle for financing development given low levels of domestic
savings and limited access to private capital flows. The amount of ODA based on UN data reaching $133 billion in 2011 (UN, 2012 P.V).

The government plays a considerable role in the planning and implementing of projects in developing countries, evidently, a large proportion of official development assistance is channelled through the public sector. The final effects of official development assistance on economic growth of the country will depend on how it influences fiscal policy (Machado, 2009).

Net official development assistance to developing countries has averaged between 3.7 and 6.7 percent of GDP during 1980–2009, amounting to around 20–40 percent of average tax revenues. The relatively high share of aid in government budgets in some countries has raised concerns about the detrimental effects of aid dependency on domestic revenue effort, spending programs and budget planning as well as institution-building (Gupta, Clements & Tareq, 2008).

Political economy considerations provide additional support to the argument that official development assistance may discourage taxation by recipient governments (Benedek, Crivelli, Gupta, & Muthoora, 2012). A key argument of the aid dependency literature, for example, is that official development assistance lowers tax revenue, because it undermines the development of domestic institutions that support tax administration and good governance (Knack, 2000; Heller and Gupta, 2002; Brautigam and Knack, 2004; and Mosley, Hudson & Horrell, 2008).

The present paper develops a fiscal response model based on the recent work in the fiscal response literature and provides solution that can be estimated using more advance econometric techniques. A model for a sample of 25 countries in Asian and Latin American countries over the period 1991–2010 have been estimated.

The paper is organized as follows. The next Section review the relevant literature, Section III outlines the specification of a fiscal response model and derive its solution to analyze the fiscal effects of official development assistance. Section IV deals with data issues. The main results will be presented in Section V and concluding remarks will be given in Section VI.

1- These countries will be mentioned in Section IV.
2- Review of literature

The fiscal response literature pioneered by Heller (1975). Fiscal behavior means deciding between various sources or revenue and level of expenditure. Specifically, it looks at how government behaves vis-à-vis of official development assistance (ODA). Critics argue that ODA have resulted in increased public or private consumption rather than in increased investment, and have contributed less to growth than was anticipated.

A data plot suggests a negative association between total net ODA and total tax revenues (Figure 1). Between 1980 and 1995, when official development assistance as a share of GDP was increasing, average tax revenue in relation to GDP decreased slightly. Post-1995, a decline in the share of total net ODA to GDP was accompanied by higher tax revenues as percentage of GDP.

![Figure1: Average Net ODA and Tax Revenue in Low- and Middle- Income Countries, 1980-2009](image)


Evidence of a negative relationship between net official development assistance and tax revenue has been borne out in the data in several studies. Ghura (1998), using data for a sample of 39 sub-Saharan African countries for the period 1985–96, finds that aid has a statistically significant negative effect on the ratio of taxes to GDP.

Gupta et al (2004) examined the revenue response to net official development assistance and found that total net ODA has a negative
association with government revenues. They also concluded that the effect depends on the composition of aid: while grants have a negative effect on revenues, loans have no significant or positive effect. They argued that their results may reflect the fact that loans have to be repaid eventually, while grants represent real transfers from abroad. Moreover, they found that although the effect of grants is modest on average, it is higher in countries with weak institutions.

Remmer (2004) uses a broader sample of 120 developing countries over the period 1970–99. She finds a negative relationship between three different measures of net official development assistance and changes in tax revenue to GDP. As noted above, Gupta et al. (2004) find similar results. Moreover, they show that the negative effect of ODA grants is stronger in countries with weak institutions.

Knack (2009) also finds a robust negative relationship between sovereign rents (from aid and natural resources) and taxation. A number of recent contributions have however argued that the results are sensitive to the composition and scope of the estimation sample.

Clist and Morrissey (2011), extending the time period of the Gupta et al. (2004) database, and re-estimating the relationship between net official development assistance and taxation from the mid-1980s to 2005 find a positive association between net official development assistance and taxation. They argue that this reflects a break in pattern of allocation of net official development assistance since the mid-1980s, when net official development assistance became conditional on implementation of good policies or concomitant structural reforms. This argument is central to Carter’s (2011) critique of existing studies on net official development assistance and taxation. He argues that the econometric evidence to date is inconclusive because of failures to adequately address the endogenous nature of net official development assistance.

for government revenue? Does it mainly finance government consumption or public investment? Does it crowds-out public borrowing?

3- A Fiscal Response Model

Fiscal response studies assume that public sector decision-makers are faced with the task of allocating resources among different alternative subject to budgetary constraints. These decision-makers in the public sector behave as a single individual with a well-behaved, homothetic preference map with following utility function following Heller (1975):

\[ U_i = F(INEX_i, CUEX_i, GREV_i, ODA_i, BORR_i) \] (1)

Where INEX is government investment expenditure, CUEX is government current expenditure, GREV is government revenue, ODA is net official development assistance and BORR is net public borrowing from other sources.

The standard approach in the fiscal response literature is to define the public sector utility function as a quadratic loss function and minimize it as follows (Ouattara, 2006):

\[ U_i = a_0 - \frac{a_1}{2}(INEX_i - INEX_i^*)^2 - \frac{a_2}{2}(CUEX_i - CUEX_i^*)^2 - \frac{a_3}{2}(GREV_i - GREV_i^*)^2 - \frac{a_4}{2}(BORR_i - BORR_i^*)^2 \] (2)

The starred variables represent exogenous target and represent the weight attached to each component of the utility function. All other variables in equation (2), represent actual variables, and treated as endogenous.

The rationale for specifying the utility function in this manner is that public decision-makers set annual targets (represented by the starred variables) for each of the decision variables and consciously strive to achieve
these targets subject to budgetary constraint\(^1\). The utility function is specified in a way that if the decision-makers try to derive from the targets it will result in a loss of utility\(^2\).

Following Ouattara (2006), the decision-maker is assumed to maximize utility function (2) subject to the following budget constraint:

\[
INEX_i + CUEX_i = GREV_i + ODA_i + BORR_i
\] (3)

The above budget constraint assumes that expenditure must equal total government's receipts. The lagrangian function is applied to maximization problem, as follows:

\[
L = a_0 - \frac{\alpha_1}{2} (INEX - INEX^*)^2 - \frac{\alpha_2}{2} (CUEX - CUEX^*)^2 - \frac{\alpha_3}{2} (GREV - GREV^*)^2 - \frac{\alpha_4}{2} (BORR - BORR^*)^2 + \lambda [INEX + COEX - GREV - ODA - BORR]
\] (4)

Where \(\lambda\) is the Lagrange multiplier? Taking the derivative of L with respect to the choice variables and \(\lambda\), and solving the first-order condition through leads to the following semi-reduced form equations:

\[
INEX = \delta_1 INEX^* + \delta_2 [ODA + BORR^* + GREV^* - CUEX^*]
\] (5)

\[
CUEX = \delta_3 CUEX^* + \delta_4 [ODA + BORR^* + GREV^* - INEX^*]
\] (6)

\[
GREV = \delta_5 GREV^* - \delta_6 [ODA + BORR^* - INEX^* - COEX^*]
\] (7)

\[
BORR = \delta_7 BORR^* - \delta_8 [ODA + GREV^* - INEX^* - CUEX^*]
\] (8)

\(^1\) Recent papers such as Franco-Rodriguez et al. (1998), McGillivray (2000) have endogenised aid on the basis that it is a government choice variable. However, given that the target for aid is generally set as aid commitments this implies that the impact of aid obtained in the reduced form equations will be that of the commitment values and not the disbursement values. This will tend to overestimate the impact of aid, as the amounts committed are generally higher than those disbursed.

\(^2\) Bihn and McGillivray (1993) provide detail discussion about the specification of the utility function.
Where the $\delta_j (j=1, 2, \ldots, 8)$, are combinations of $\alpha_i (i=1, \ldots 4)$

One problem faced with most fiscal response studies has been to obtain the target variables. It is assumed that the target for public capital expenditure can be represented by the following:

$$INEX^*_i = \gamma_0 + \gamma_1 GDP_{c_i} + \gamma_2 DESE_i + \gamma_3 ODA_i$$ \hspace{1cm} (9)

Where GDP is GDP per capita, DESE is debt service and ODA is net official development assistance. GDP measures the country's level of economic development. Public capital expenditure will also depend on the debt burden of the country measured here by debt servicing. Target for government consumption expenditure is also defined by a similar economic relationship:

$$CUEX^*_i = \eta_0 + \eta_1 GDP_{c_i} + DESE_i + \eta_3 ODA_i$$ \hspace{1cm} (10)

The target for government revenue is approximated by:

$$GREV^*_i = \mu_0 + \mu_1 GDP_{c_i} + \mu_2 EXPR_i + \mu_3 IMPR_i$$ \hspace{1cm} (11)

GDP is defined as above, EXPR and IMPR denotes exports and imports respectively. The inclusion of EXPR and IMPR is based on the fact that many developing countries get revenue from exports and imports. Finally, the target for borrowing is specified by the following relationship:

$$BORR^*_i = \varepsilon_0 + \varepsilon_1 GDP_{c_i} + \varepsilon_2 ODA_i$$ \hspace{1cm} (12)

Borrowing will depend on GDP and ODA. Substituting Eq. (9)-(12) into Eq. (5)-(8) gives the following full reduced form equations:

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1- These coefficients can be obtained from the authors upon request.
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\begin{align*}
\text{INEX}_{it} &= \omega_0 + \omega_1 \text{GDP}_{c1it} + \omega_2 \text{DESE}_{it} + \omega_3 \text{ODA}_{it} + \omega_4 \text{EXPR}_{it} + \omega_5 \text{IMPR}_{it} + \nu_{1i} + \epsilon_{1it} \\
\text{CUEX}_{it} &= \psi_0 + \psi_1 \text{GDP}_{c1it} + \psi_2 \text{DESE}_{it} + \psi_3 \text{ODA}_{it} + \psi_4 \text{EXPR}_{it} + \psi_5 \text{IMPR}_{it} + \nu_{2i} + \epsilon_{2it} \\
\text{GREV}_{it} &= \pi_0 + \pi_1 \text{GDP}_{c1it} + \pi_2 \text{DESE}_{it} + \pi_3 \text{ODA}_{it} + \pi_4 \text{EXPR}_{it} + \pi_5 \text{IMPR}_{it} + \nu_{3i} + \epsilon_{3it} \\
\text{BORR}_{it} &= \rho_0 + \rho_1 \text{GDP}_{c1it} + \rho_2 \text{DESE}_{it} + \rho_3 \text{ODA}_{it} + \rho_4 \text{EXPR}_{it} + \rho_5 \text{IMPR}_{it} + \nu_{4i} + \epsilon_{4it}
\end{align*}

\(\omega\)'s, \(\psi\)'s, \(\pi\)'s and \(\rho\)'s can be traced back to the \(\alpha\)'s representing the weigh attached to each element of utility function (2). In this regard panel data techniques provide us with more powerful tools to test these relationships among economic variables.

In this paper, the above equations will be estimated using panel data econometrics technique. Given that equations derived from the solution of the model. For this purpose each of Eqs. (13)- (16) can be represented by the following general reduced-form regression:

\[Y_{it} = \alpha + \beta X_{it} + e_{it}\]  

Where the \(Y_{it}\) is dependent variable and \(X_{it}\) denotes a set of explanatory variable for country \(i\) in time \(t\).

These two types of estimators are designed to handle the systematic tendency of \(e_{it}\) to be higher for some individuals than for others (individual effects) and possibly higher for some time periods than for others (time effects). The fixed effect estimator does this by (in effect) using a separate intercept for each individual or time period. The random effect estimator is based on the following decomposition of \(e_{it}\) where \(e\) is the individual effect, \(\mu\) the time effect, and \(\eta\) the purely random effect. \(\beta\) is estimated by the structure imposed upon \(e_{it}\) by this assumption. Both fixed and random effects adjust for heteroskedasticity. However, the two effects have their own advantages and disadvantages. The random effects model treats the
individual effect as part of the error term, it suffers from the possibility of bias due to a correlation between it and the repressors.

Hausman (1978) provides a test for discriminating between the fixed effects and the random effects estimators. The test is based on comparing the difference between the two estimators of the coefficient vectors, where the random effects estimator is efficient and consistent under the null hypothesis and inconsistent under the alternative hypothesis. The fixed effects estimator is consistent under both the null and the alternative hypothesis. If the null is true then the difference between the estimators should be close to zero. The calculation of test statistics requires the computation of the covariance matrix of $\beta_1$ - $\beta_2$. In the limit the covariance matrix simplifies to $\text{Var}(\beta_1) - \text{Var}(\beta_2)$, where $\beta_1$ is the fixed effects estimator and $\beta_2$ is the random effects estimator.

4- Sources of Data

The data used here come from two main sources and cover the period 1991-2010. Countries in Southland West Asia are: Iran, Turkey, Pakistan, Jordan, Oman, Syria, India, and Yemen. In South and East Asia are: Indonesia, Malaysia, Thailand, Philippine, Bangladesh, Nepal, Vietnam. In Latin America are: Argentina, Brazil, Bolivia, Venezuela, Colombia, Peru, Chile, Paraguay, Uruguay. Data on government investment expenditure is obtained from the World Bank Global Development Network Growth Database. Data on government current expenditure, government revenue, foreign aid, debt service, exports and imports of goods and services (all expressed as a percentage of GDP), GDP$_C$ is GDP (market price) per capita were obtained from the World Bank World Development Indicator, 2012. Data on borrowing is the net public borrowing, obtained as a residual from Eq. (3) in the model.

5- Empirical Results

Tables 1-4, summarize the results related to the estimation of Eqs.(13)-(16). Each equation applied for the same period (1991-2010). The sizes of the samples are 25 countries in Asian and Latin American countries. According to F-Lemer test, $H_0$ hypothesis (Pooling approach) rejected and all equations estimated with panel data approach. According to Husman test

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1- Because other countries in these regions are not reported variables such as debt service, ODA, … are not considered.
in government investment expenditure equation, $H_0$ hypothesis accepted and this equation estimated based on random effects estimator. In other equations, $H_0$ hypothesis rejected and these equations estimated based on fixed effects estimator.

The impact of net official development assistance on the government investment expenditure is assessed in Table 1.

Table 1: Estimated Results for Government Investment Expenditure Equation, Based on Random Effects

<table>
<thead>
<tr>
<th>Variable</th>
<th>Coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td>GDPc</td>
<td>0.00027 (1.81)**</td>
</tr>
<tr>
<td>DESE</td>
<td>-0.00024 (-0.11)</td>
</tr>
<tr>
<td>ODA</td>
<td>0.170 (1.99)**</td>
</tr>
<tr>
<td>EXPR</td>
<td>0.0154 (0.61)</td>
</tr>
<tr>
<td>IMPR</td>
<td>0.0734 (2.51)**</td>
</tr>
<tr>
<td>Constant</td>
<td>5.23 (4.17)**</td>
</tr>
<tr>
<td>F- Lemer</td>
<td>50.62 (0.000)</td>
</tr>
<tr>
<td>Husman-test</td>
<td>5.23 (0.39)**</td>
</tr>
<tr>
<td>Observations</td>
<td>500</td>
</tr>
</tbody>
</table>

Number of countries 25

Robust z-statistics in parentheses. ** significant at 5%; *** significant at 1% & **** significant at 10%." Indicate P-value for F and Husman test, respectively.

Source: Authors

The results in Table 1 show that net official development assistance bears a positive and statistically significant sign in government investment expenditure. Which is consistent with findings of Heller (1975), Khan and Hoshino (1992), Franco-Rodriguez et al. (1998), Ouattara (2006), Machado (2010). The impact of foreign aid on the government current expenditure is assessed in Table 2.

Table 2: Estimated Results for Government Current Expenditure Equation, Based on Fixed Effects

<table>
<thead>
<tr>
<th>Variable</th>
<th>Coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td>GDPc</td>
<td>-0.0001 (-0.82)</td>
</tr>
<tr>
<td>DESE</td>
<td>0.011 (0.30)</td>
</tr>
<tr>
<td>ODA</td>
<td>0.047 (0.80)</td>
</tr>
<tr>
<td>EXPR</td>
<td>-0.041 (-2.43)**</td>
</tr>
<tr>
<td>IMPR</td>
<td>0.069 (3.41)***</td>
</tr>
<tr>
<td>Constant</td>
<td>10.99 (20.45)**</td>
</tr>
<tr>
<td>F- Lemer</td>
<td>57.25 (0.000)</td>
</tr>
<tr>
<td>Husman-test</td>
<td>18.25 (0.000)</td>
</tr>
<tr>
<td>Observations</td>
<td>500</td>
</tr>
</tbody>
</table>

Number of countries 25

Robust t-statistics in parentheses. ** significant at 5%; *** significant at 1%." Indicate P-value for F and Husman test, respectively.

Source: Authors
The results in Table 2 show that net official development assistance has insignificant effect on government current expenditures, then contrary to the finding of Boone (1996) and others, we find no evidence that aid flows are used to finance government consumption expenditure. The impact of foreign aid on the government revenue is examined in Table 3.

Table 3: Estimated Results for Government Revenue Equation, Based on Fixed Effects

<table>
<thead>
<tr>
<th>Variable</th>
<th>Coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td>GDPC</td>
<td>-0.0002 (-1.75)**</td>
</tr>
<tr>
<td>DESE</td>
<td>-0.063 (-1.20)</td>
</tr>
<tr>
<td>ODA</td>
<td>-0.185 (-2.25)**</td>
</tr>
<tr>
<td>EXPR</td>
<td>0.019 (0.80)</td>
</tr>
<tr>
<td>IMPR</td>
<td>0.20 (0.71)</td>
</tr>
<tr>
<td>Constant</td>
<td>19.23 (25.56)**</td>
</tr>
<tr>
<td>F- Lemer</td>
<td>57.25 (0.000)+</td>
</tr>
<tr>
<td>Husman-test</td>
<td>17.82 (0.003)+</td>
</tr>
</tbody>
</table>

Robust t-statistics in parentheses. ** significant at 5%; *** significant at 1% & **** significant at 10%. +Indicate P-value for F and Husman test, respectively.

Source: Authors

The coefficient of net official development assistance on revenue is negative and statistically significant at the conventional level, thus suggesting that developing countries public sector authority reduce their revenue collection efforts when net official development assistance is made available to them. This finding corroborates finding by Heller (1975), Franco-Rodriguez et al. (1998), Mc- Gillivray, Ouattara (2005). Finally and Benedek et al (2012). Table 4 which summarize results related to the net public borrowing in equation (16).

Table 4: Estimated Results for Public Borrowing Equation, Based on Fixed Effects

<table>
<thead>
<tr>
<th>Variable</th>
<th>Coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td>GDPC</td>
<td>0.0004 (2.88)**</td>
</tr>
<tr>
<td>DESE</td>
<td>0.065 (1.18)</td>
</tr>
<tr>
<td>ODA</td>
<td>-0.56 (-6.61)**</td>
</tr>
<tr>
<td>EXPR</td>
<td>-0.043 (-1.74)**</td>
</tr>
<tr>
<td>IMPR</td>
<td>0.13 (4.51)**</td>
</tr>
<tr>
<td>Constant</td>
<td>-3.58 (-4.52)**</td>
</tr>
<tr>
<td>F- Lemer</td>
<td>30.94 (0.000)+</td>
</tr>
<tr>
<td>Husman-test</td>
<td>20.14 (0.001)+</td>
</tr>
</tbody>
</table>

Robust t-statistics in parentheses. ** significant at 5%; *** significant at 1%. +Indicate P-value for F and Husman test, respectively.

Source: Authors
The results shows that net official development assistance have a negative and statistically significant effect on public sector borrowing. Overall this results support the view that government in developing countries are likely to reduce other forms of borrowing.  

6- Concluding Remarks

This paper has investigated the relation between net official development assistance and public sector fiscal behavior using a panel data in Asian and Latin American countries over the period 1991-2010. The results indicate that net official development assistance has a positive and significant effect on investment expenditure, which in turn might crowds in private investment. The evidence also shows that ODAloans are not used to increase current expenditure. Results also show that net official development assistance crowds out government revenue, then public sector authorities in developing countries reduce their efforts to collect revenue when ODA is available to them. Finally, net official development assistance does induce a reduction in public sector borrowing.

This paper has also some research implications. First, as argued by McGillivray and Morrissey (2001) ODA effectiveness studies need to first consider the question of government behavior vis-à-vis of ODA loans before studying their broad macroeconomic impacts. The relationship between ODA and growth might not be a straightforward one, as many aid growth studies seem to assume. Therefore, understanding how these ODA loans affect fiscal aggregates might shed new light on the aid debate. Second, it might also be worth analyzing whether the fiscal behavior of the recipient government depends on the nature or the type of ODA loans.

This implication of our findings is that developing country should attract more foreign capital by liberalizing their financed markets.

References


1 - The argument is based on the fact that aid is cheaper compared to borrowing.


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