

The Monetary Policy, Credit Constraint and Spatial Distribution of Economic Activity: A Contribution of New Economic Geography*

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Abstract

This paper investigates the effect of monetary policy on the distribution of economic activity and agglomeration economies within a country. The considered channel for this effectiveness is the availability of credit to firms in various regions and the effects on the labor and consumer welfare. For this purpose, data for manufacturing firms located in 30 different provinces in Iran during 2007 and 2014 gathered. The empirical results from spatial panel data show that beside conventional channel of effectiveness through consumer and labor force utility function, regional monetary policy implication through uneven distribution of regional loanable banking fund seems to be substantial centripetal force. In terms of most well-known NEG variable, uneven regional accessibility of credit market has opposite regional implication as trade freeness. While the former leads to more concentration of economic activity across space, the latter tends to drive dispersion. It is assumable that monetary policy reduce the impact of credit constraints on firms but the degree of credit availability in regions is a significant driver for concentration of economic activity. The result shows the importance of accessibility to banking loans on distribution of economic activities within the country.

Keywords: Monetary Policy, Credit Constraint, Agglomeration, Dispersion, Spatial Panel Data, Dynamic Panel Data, Iran.

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1. Introduction

Empirical studies shows that there is an inherent differences in the regional wages in a growing number of countries (Duranton and

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Monastiriotis, 2002; Azzoni and Servo, 2002; Maier and Weiss, 1986). Identical workers in terms of various measure of human capital such as education and experience are paid differently for identical work depending upon the region in which they work (Beenstock and Felsenstein, 2008). However, some studies suggest that just as there is convergence failure at the international level (Barro et al., 1991), there is also convergence failure within countries (Beenstock and Felsenstein, 2010). It seems surprising as trade is likely to be freer within countries than between them and capital and labor force which assume has the same characteristic are likely to be more mobile within countries.

To explain the discovery of convergence failure at the international level new theories, such as endogenous growth theory (e.g. Grossman and Helpman, 1993) was generated. To explain the fact of regional convergence failure the same has happened and new theory such as new economic geography (NEG) try to answer this surprise (Beenstock and Felsenstein, 2010).

Initially originated from international trade theory, New Economic Geography (NEG), is an attempt to find factors which shape firms and workers' location behavior and to explain the formation of a large variety of economic agglomeration (or concentration) in geographical space (Fujita and Krugman, 2004).

Even though the traditional location choice factors have been changing in time, initial advantages, low transportation costs, accessibility to market and skilled labor, are still significant forces generating agglomeration economies (Fujita and Thisse, 1996; McCann, 2001; Parr, 2002; Capello, 2007).

The economics of agglomeration, whose origins can be traced back to the work of Marshall (1898; 1919; 1930), basically tend to be categorized into kinds of external economies – a pooled specialized labor market, specialized factor of production suppliers and technological spillovers (Artis et al., 2011). Technological spillovers as the third part of Marshallian agglomeration theory consist of informational or knowledge externalities which result from the concentration of (both vertically and horizontally) related firms, facilitating processes of learning and innovation in the locality (Malmberg and Maskell, 1997; 2002).

According to the location patterns of manufacturing industry, on one hand firms are likely to concentrate within the metropolitan areas when they have access to a larger markets and lower transport costs and on the other hand, urban areas provide a wide array of final goods and specialized labor market pool which make them attractive to consumers and workers. As a result of the process, Agglomeration economies are the result of cumulative processes involving both the supply and demand sides (Krugman, 1980; 1991; Ottaviano and Thisse, 2003; Puga, 2010).

In this paper we examine the effect of monetary policy on the distribution of economic activity within the country. For this purpose, data for manufacturing firms located in 30 different provinces of Iran during 2007-2014 are gathered. The innovative approach of this study is regarding to inclusion of monetary policy into NEG models, which basically are silent toward this kind of macroeconomic policy. The way of effectiveness of monetary policy is loanable banking resource in each regions and though the availability of credit to firms in various region. High degree of spatial concentration in the financial market in developed and metropolitan areas (Palmberg, 2012) faced firms in less developed areas with critical performance challenges as informational disadvantages (Danielson and Scott; 2004, Petersen and Rajan; 1994, 1997), limited source and higher cost of borrowing (Arena and Dewally, 2012; Smith, 1987; Petersen and Rajan, 1994), higher risk of asset substitutions (Leland and Pyle, 1977), and inflexibility of capital structure, sub-optimally lower leverage ratio (Arena and Dewally, 2012; Mayers, 1977), lead them to have a poorer performance and will change to a significant obstacle to their expansion, join new markets and export orientation production (Fauceglia, 2015). Due to such critical obstacles unevenness distribution of regional loanable banking fund seems to be substantial centripetal force.

It is assumable that monetary policy expansion reduce the impact of credit constraints on firms (Orlowski, 2015) but the degree of credit availability in regions is a significant driver for concentration of economic activity. Besides monetary policy has a direct effect on the labor force and consumer utility which affect the migration decision.

The outline of the rest of the paper is as follows. Section 2 review the relevant literature of financial market and banking performance on

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the economic agglomeration and provides a framework to analyze the effect of monetary policy on the spatial dispersion and concentration. In the section 3 explanation of econometric model, variables, data and empirical result are presented. Finally, last section devoted to conclusion and policy implications.

2. Literature

Although financial economics literature provides limited insight on the relationship between firm's geographical location and capital structure and credit accessibility (Arena and Dewally, 2012), empirical evidence shows that there is a significant interplay among geographical location and structure and amount of financial funds available to firms (Coval and Moskowitz, 1999; Grinblatt and Keloharju, 2001; Loughran and Schultz, 2005; Malloy, 2005; Loughran, 2008). Some studies argue that firms in regions with poorly developed financial institutions as they cannot borrow sufficiently from financial institutions when they require external finance they face tough operational problem and though have to find other source like receive more trade credit (Ge and Qiu, 2007; Fisman and Love, 2003). This will be more severe for small businesses as banks do not typically offer them sufficient credit because of the presence of information asymmetry (Danielson and Scott; 2004, Petersen and Rajan; 1994, 1997). Informational disadvantage faced by firms located in less developed and rural area is a significant factor to increases their cost of borrowing and prevent them to establish a sound financial relationship with a large number of lending institutions (Arena and Dewally, 2012; Brickley et al., 2003). In such cases negotiation, relationship between borrower and lender are usually based on imprecise or soft information which cannot be verifiable through official documents (Berger et al., 2005). Debt-holders might impose higher yields on rural firms to compensate for weak information and higher risk of asset substitution (Leland and Pyle, 1977). Besides, several studies shows that banks are able to receive more in depth information about a borrowing firm's quality when the geographical distance between the lending institution and the borrower is shorter (Dass and Massa, 2011; Hauswald and Marquez, 2006). This may lead to squeeze the source of funding for rural and small cities firm to

borrow more from local banks and repeatedly borrow from the same banks instead of frequently changing lenders (Arena and Dewally, 2012). In additions instead of short-term debt, to raise debt capital banks might be more likely to offer small area firms longer-term debt which might increase the underinvestment problem and sub-optimally lower leverage than urban firms (Arena and Dewally, 2012; Mayers, 1977). During monetary contraction availability of credit has more detrimental effect on firm's performance (Atanasova and Wilson, 2004; Choi and Kim, 2005 and Mateut et al., 2006). Besides firms with high levels of short-term debt, which are vulnerable to financial crises, reduce the provision of trade credit during periods of contraction in bank credit which may be highly risky (Love et al., 2007). There is some evidence shows that small firms did not receive any alternative source of sufficient credit to compensate for the decline in bank loans during a monetary squeeze (Marotta, 1997). By focusing on leverage and debt maturity there is expressive interplay between geographical location of firms and the structure of corporate debt. Empirical finding of the effect of proximity on firms equity shows that firms in less developed area cannot easily change their capital structure from debt to equity (or vice versa) to reduce their cost of capital (Loughran, 2008).

There is high degree of spatial concentration in the financial market in developed and metropolitan areas which illustrates the importance of local embeddedness, networks, tacit knowledge and face-to-face communication, knowledge spillovers, and spatial proximity for the organization of the industry (Palmberg, 2012). Also there is a general consensus that banking system just same as equity institutional investors, debt investors, debt underwriters are clustered in urban and developed areas. In such circumstances it is not surprising that as the distance between borrowing firms and their banks is increasing, bank lending is still principally tend to be local (Petersen and Rajan, 2002; Becker, 2007). Consequently small city firms and rural are more likely to rely on local banking system which face squeeze loanable resources. Alternative firm's compensation of reduced availability of bank loans sometimes involving annual interest rates in excess of 40% (Smith, 1987; Petersen and Rajan, 1994). Beside such expensive financial cost, to expand performance they cannot offer more trade

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credit to their customer (Tsuruta, 2014; Montoriol-Garriga, 2013) and not able to join new markets and export orientation production (Fauceglia, 2015).

Krugman and Venables (1990) provide one of the first relevant formal contributions about the functioning of NEG models in predicting agglomeration in a framework of economic integration (Ascani et al., 2012). The core-periphery model has two main variants. The footloose-labor variant (Krugman, 1991) and the vertical-linked-industries variant (Venables, 1996). In the first one agglomeration forces driven by inter-regional labor migration within a single sector and the motivation of the migration is the differences in the regional real wage. In the latter one agglomeration driven by intersectoral migration within each region; and the intersectoral nominal wage differences motivate migration (Baldwin, 2001).

Fujita et al. (1999) –FKV–provides a comprehensive review on the standard CP model. In the initial stage of CP model there is two symmetric region, two factors of production and two sector of manufactures and agriculture. Manufacture sector is a Dixit-Stiglitz monopolistic competition with increasing returns whereas the agriculture sector has a perfectly competitive production function with constant return which produce homogenous good. Production in both sector is tradable but in the monopolistic competition there is a fractional trade cost which assumed as iceberg trade cost and in the perfectly competitive sector is costless.

Dixit-Stiglitz monopolistic competition emerges as a market structure determined both by consumers' preferences for variety and firms' fixed requirements for limited productive resources. Description of such quality is on the demand side, preference of consumers for variety and on the supply side, internal economies of scale for each good, but no economies of scope across goods (Fujita and Thisse, 2009).

Representative consumer preferences is a Cobb-Douglas function of the consumption of agriculture and manufacture production:

$$U \equiv C_x^\mu C_z^{1-\mu}; \quad C_x \equiv \left(\int_{i=0}^{n+n^*} c_i^{1-\frac{1}{\sigma}} \right)^{\frac{1}{1-\frac{1}{\sigma}}}; \quad 0 < \mu < 1 < \sigma \quad (1)$$

Where C_x represents a composite index of the consumption of manufacture good and C_z is the consumption of agriculture good. μ is a constant representing the expenditure share of manufactured good, n and n^* are the number (mass) of varieties in two regions, and σ represents the elasticity of substitution between any two varieties of manufactured goods.

Regional supplies of agriculturist (A) as well as the global supply of workers (L) are fixed, but the inter-regional distribution of L is endogenous and determined by regional real wage differences.

Assuming s_L the share of labour in the north, L and L^w are the north work supply and total work supply, ω , ω^* , $\bar{\omega}$ are the northern, southern and average real wages. P is the north region price index with p_z being the price of manufactured product and p_i being the price of variety i ; the exposition of migration equation in the KfV model has showed as below;

$$\dot{s}_L = (\omega - \bar{\omega})s_L; \quad s_L \equiv \frac{L}{L^w}; \quad \omega = \frac{W}{P}; \quad \bar{\omega} \equiv s_L\omega + (1 - s_L)\omega^* \quad (2)$$

$$P \equiv p_z^{1-\mu} \left(\int_{i=0}^{n+n^*} p_i^{1-\sigma} di \right)^{\mu/(1-\sigma)} \quad (3)$$

By optimization of Eq.1 a constant division of expenditure between manufactured goods (X) and agriculture good (Z), CES demand functions for manufactured good varieties, agricultural good and expenditure function (E) can be written as:

$$c_j = \frac{p_j^{-\sigma} \mu E}{\int_{i=0}^{n+n^*} p_i^{1-\sigma} di}; \quad (4)$$

$$C_z = \frac{(1-\mu)E}{p_x} \quad (5)$$

$$E = wL + w_A A \quad (6)$$

The Eq. 6 is a good expression for demand-linked circular causality or backward linkages as an agglomeration force in the model. From this equation migrants can be viewed as consumers. Starting from

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symmetry, a small displacement between two regions may change the size of the market in both region. By changing the size of the market, firms sales in north raise and fall in the south. This encourages northern firms to hire workers, southern firms to fire workers, and thus the small migration shock encourages more migration (Baldwin, 2001).

On the supply side, since A is immobile and both region produce some agricultural good, free trade in agricultural good equalizes wage rates, w and w^* in two regions. Regarding price of manufactured and agricultural products by choosing units of agricultural product such that one unit of agriculturist (A) is required per unit of agricultural product, $p_z = w = w^* = 1$. Also measuring manufactured products in units such that $a_x = (1 - \frac{1}{\sigma})$ the price of a northern manufactured product variety in its local and export markets are $p = w$ and $p^* = w\tau$ respectively. Also we have;

$$\pi = \frac{w(L-nF)}{(\sigma-1)n} \quad (7)$$

The free entry condition requires n to rise to the point where $\pi = wF$. Using Eq. 7 we have;

$$n = \frac{L}{\sigma F} ; \bar{x} = \sigma F \quad (8)$$

Where \bar{x} is the equilibrium firm size of a typical firm in manufacturing sector. Eq. (8) serve as the second agglomeration force in the model which views migrants as workers. An exogenous increase in L and corresponding decrease in L^* would raise n and lower n^* . Since locally produced varieties attract no trade cost the shift in n 's would, other things equal (in particular the w 's), raise the north's relative real wage, This in turn would tend to pull in more migrants. This is called cost-linked circular causality, or forward linkages (Baldwin, 2001).

By introducing the credit constraints we can assess the effect of monetary policy on the firm's production decision. Due to level of productivity φ and internal funds ω , a firm decides simultaneously

whether to supply the home region and export to the other regions. Matsuyama (2005) introduced the effect of credit constraints in a general way. In this framework as a consequence of imperfections in financial contracting firms can only borrow a fraction (θ) of their operating profit which it correspond to maximum amount to cover fixed production expenditure and exporting cost, f and f_x , though internal funds play crucial role to finance remaining part.

An exporting firm must consider an additional fixed exporting cost f_x and an iceberg trade cost τ , where $\tau > 1$ of each good must be shipped in order for one good to reach the export destination. It is assumable regarding Melitz (2003) type models that $\tau^{\sigma-1}f_x > f$ and the cut-off productivity level for exporting profitably φ_x is higher than the productivity threshold to earn nonnegative profits in the domestic market φ^* .

For exploiting greater market firms should overcome potential financing obstacles. Only firms that meet the following export profitability condition (9) and the credit constraint condition (10) will therefore become exporters and can gain from grater market:

$$\tau^{1-\sigma} \frac{\mu}{\sigma} E(P\rho)^{\sigma-1} (\varphi)^{\sigma-1} \geq f_x \quad (9)$$

$$\theta \left[\frac{1}{\sigma} (1 + \tau^{1-\sigma}) r_d(\varphi) \right] \geq f + f_x - \omega \quad (10)$$

Accruing positive export gain due to the export profitability condition (9) need minimum level of exporter productivity of $\varphi_x = \frac{\tau}{P\rho} \left(\frac{\sigma f_x}{\mu E} \right)^{\frac{1}{\sigma-1}}$. In additions, neutralizing credit constraint condition (10), granting external finance and availability of internal funds yields the minimum cutoff firm productivity of;

$$\bar{\varphi}_x(\omega, \theta) = \frac{1}{P\rho} \left(\frac{\sigma(f+f_x-\omega)}{\theta \mu E (1+\tau^{1-\sigma})} \right)^{\frac{1}{\sigma-1}} \quad (11)$$

Only firms that draw a firm productivity at least as high as $\varphi \geq \max[\varphi_x, \bar{\varphi}_x(\omega, \theta)]$ are able to export profitably and secure access to finance (Fauceglia, 2015).

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Eq (4) to (6) and (8) gives the market clearing condition as below;

$$w\bar{x} = R; \quad R \equiv \frac{w^{1-\sigma}\mu E}{nw^{1-\sigma} + \phi n^* w^{*1-\sigma}} + \frac{\phi w^{1-\sigma}\mu E^*}{\phi n w^{1-\sigma} + n^* w^{*1-\sigma}} \quad (12)$$

Where R is a mnemonic for ‘retail sales’ and $\phi = \tau^{1-\sigma}$ measures ‘free-ness’ of trade. Variation of the free-ness of trade rises from $\phi = 0$ which means infinite trade costs to $\phi = 1$ which mean zero trade costs. Eq (12) serve as stabilizing and dispersion force in the model. By moving a small mass of L from south to north and raise n and lower n*, from the expression for R, this tends to increase competition for consumers among northern firms, thus lowering R. Though northern firms would have to pay a lower nominal wage. Consequently the drop in w and corresponding rise in w* would make north less attractive to workers. In the core-periphery literature, this dispersion force is commonly called the ‘local competition’ effect or ‘market-crowding’ effect (Baldwin, 2001).

The relation between the level of trade costs and agglomeration and dispersion forces can be conveniently summarized by Fig. 1. Indeed, as pointed out by Baldwin et al., 2003, the scenario depicted by Fig. 1 is broadly consistent with most NEG models, both static (e.g. Krugman 1991a; Krugman and Venables, 1995; Venables, 1996; Puga, 1999; Ottaviano et al., 2002) and dynamic ones (e.g. Baldwin, 1999; Martin and Ottaviano, 1999, 2001; Baldwin et al., 2001).

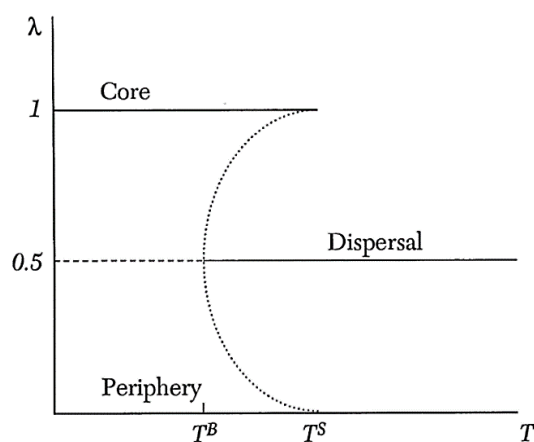


Figure 1: Agglomeration and Dispersion Equilibria as a Function of Trade Costs

Figure 1 portrays the possible long-run spatial configurations of a simple economy consisting of two regions with no inner spatial dimensions. The figure illustrates how the number and type of equilibria vary with the level of trade cost T . The vertical axis measures λ , the share of firms located in one region, solid and dotted lines denote stable and unstable equilibria respectively. At every level of trade costs there exists a symmetric diversified equilibrium (Neary, 2001). In the figure the extent of trade freeness, T is represented on the horizontal axis while the share of firms located in one of the regions appears on the vertical one. Trade freeness is an inverse measure of trade costs: $T = 1$ means autarky; $T = 0$ means free trade. Heavy solid lines indicate long-run outcomes. These are geographical distributions of firms towards which the economic system evolves as pointed out by the vertical arrows. Fig. 1 then shows that for low trade freeness ($T > T^S$) a dispersed geographical distribution of firms is the only long-run outcome. For high trade freeness ($T < T^B$) agglomeration in either region is the only long-run outcome. For intermediate values of trade freeness ($T^B < T < T^S$) both dispersion and agglomeration can emerge in the long run, so history and policy have a potential role in influencing which equilibrium prevails (Ottaviano, 2003; Neary, 2001).

What should be emphasized is that new economic geography theory does leave space for other factors such as economic policies and geography to play their roles. As stated by Neary (2001), when trade costs are in certain range, both agglomeration and diversification are possible equilibria, so history and policy have a potential role in influencing which equilibrium prevails.

3. Econometric Model and Data Explanation

The model adopts the following form:

$$AG_{it} = \beta_0 AG_{i,t-1} + \beta_1 CR_{it} + \beta_2 W_{it} + \beta_3 CPI_{it} + \beta_4 TRV_{it} + \beta_5 GDPC_{it} + \beta_6 HC_{it} + \beta_7 GB_{it} + \beta_8 UD_{it} + \varepsilon_{it} \quad (13)$$

Where (AG) is the various index of agglomeration, (CR) monetary policy stance and availability of credit, (W) regional manufacturing

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wage rate, (*CPI*) regional consumer price index, (*TRV*) regional share of transport and communication value added as an index of freeness of trade, (*GDPC*) regional GDP per capita to capture the market size effect, (*HC*) regional human capital quality, (*GB*) regional government budget and (*UD*) is an index for urban development.

Economic policies have their spatial impacts. Particularly, we will test the impact of two types of policies on industrial agglomeration, monetary policy and the government involvement in regional economic activities. Monetary policy is represented by credit available in each region. Obviously, different credit availability in various regions is expected to encourage regional industrial agglomeration. To investigate the local government involvement in regional economic activities, regional government expenditure included in the model which we expect to weaken the regional industrial agglomeration. Lag independent variables are used as proxy variables which shows the importance of history and previous industrial structure and shows the effect of history and geography on regional industrial agglomeration (Chen et al., 2008).

The new economic geography theory in the trade-off between centrifugal and centripetal forces by confirmation of existence of externality based on industrial backward and forward linkages, human capital accumulation (Henderson, 1974) and “home market effect” (Fujita 1988; Krugman, 1991) have a critical point of view to neoclassical economics. In order to test these factors that all base on increasing returns, we include the following variables into the econometric model: (i) The EG index, regional share in industry and manufacturing value added, which measures relative industrial externality; (ii) The regional literacy rate as a proxy for regional comparative advantage in human capital. (iii) The regional per capita GDP which measures the relative capacity of local market; (iv) Urban development index as the ratio of share of nonagricultural population to its national average which we think better represents the regional infrastructure. (v) Transaction cost as the ratio of the share of regional transportation, post, storage and telecommunication in GDP to the national average, which captures development of information and communication service. Since lower transaction cost helps attract

firms, this variables should be positively related to regional share in industrial activity.

Table 1: Summary of Variable and Indices

Row	Abbr	variable	Index
1	CR	Regional Monetary policy	Total paid loan of banking system in region
2	GB	Regional Government interference	Regional government budget
3	TRV	Transportation cost	Regional share of transportation, post and telecommunication in GDP
3	AG	Concentration& dispersion of economic activity	Regional EG index, regional manufacturing and industry value added
4	W&CPI	Welfare and cost of living	Regional manufacturing wage rate and regional consumer price index
5	GDPC	Home market effect	Regional GDP per capita
6	HC	Human capital development	Regional rate of literacy
7	UD	Urban development	Regional share of nonagricultural population to the total population

4. Estimation and results:

In estimating equation (13), the disturbance vector is assumed to have random region effects as well as spatially auto-correlated residual disturbances;

$$\varepsilon_t = \mu + \zeta_t \quad (14)$$

$$\zeta_t = \rho W \zeta_t + \eta_t \quad (15)$$

where $\varepsilon_t = (\varepsilon_{1t}, \dots, \varepsilon_{Nt})'$, $\zeta_t = (\zeta_{1t}, \dots, \zeta_{Nt})'$ and $\mu = (\mu_t, \dots, \mu_N)'$ denotes the vector of random region effects, which are assumed to be i.i.d. $(0, \sigma_\eta^2)$. ρ is the scalar spatial autoregressive coefficient with $|\rho| < 1$. W is a known $N \times N$ spatial weights matrix where diagonal elements are zero. In this study, the weights matrix is constructed so that a neighboring region takes the value of 1 and 0 otherwise. Rewriting (15) as:

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$$\xi_t = (I_N - \rho W)^{-1} \xi_t = A^{-1} \eta_t \quad (16)$$

Where $A = I_N - \rho W$ and I_N is an identity matrix of dimension N. Also by rewriting (14) into vector form we have:

$$\varepsilon = (\iota_T \otimes I_N) \mu + (I_T \otimes A^{-1}) \eta \quad (17)$$

Where ι_T is a vector of ones of dimension T and I_T is an identity matrix of dimension T.

The variance-covariance matrix of ε is as follow:

$$\Omega_\varepsilon = E[\varepsilon \varepsilon'] = \sigma_\mu^2 (J_T \otimes I_N) + \sigma_\eta^2 [I_T \otimes (A' A)^{-1}] \quad (18)$$

Where J_T is a matrix of one of dimension T. Following Baltagi, Song, and Koh (2003), this variance–covariance matrix can be rewritten in such a way that

$$\Omega_\varepsilon = \sigma_\eta^2 \{ \bar{J}_T \otimes [T \phi I_N + (A' A)^{-1}] + E_T \otimes (A' A)^{-1} \} = \sigma_\eta^2 \Sigma_\varepsilon \quad (19)$$

Under the assumption of normality, the log-likelihood for our model, conditional on δ , becomes (Baltagi, Song, and Koh 2003):

$$\begin{aligned} \mathcal{L}(\gamma, \sigma_\eta^2, \phi, \rho | \delta) = & -\frac{NT}{2} \ln(2\pi\sigma_\eta^2) - \frac{1}{2} \ln | \Sigma_\varepsilon | - \frac{1}{2\sigma_\eta^2} e' \Sigma_\varepsilon^{-1} e = \\ & \frac{NT}{2} \ln(2\pi\sigma_\eta^2) - \frac{1}{2} \ln | T \phi I_N + (A' A)^{-1} | + \frac{T-1}{2} \ln | A' A | - \frac{1}{2\sigma_\eta^2} e' \Sigma_\varepsilon^{-1} e \end{aligned} \quad (20)$$

According to the Hausman test, which is used for deciding whether the fixed or the random effect spatial lag model should be used, the fixed effects model is convenient for the current situation. According to the similar Hausman statistics for the fixed or random effect spatial error models, the fixed effect spatial lag model turns out to be superior. For the fixed effect specifications Spatial fixed effects lag model versus Spatial fixed effects error model, the LM statistics may be applied as indication of which type of spatial dependence should apply, It turns out that the spatial lag model is the most convenient for the present data since LM_ρ is more significant than LM_λ .

Table 2: Spatial Panel Data and Dynamic Panel Estimation:

Methods	Spatial Error Panel Data estimation			GMM Panel Data estimation		
	model 1*	model 2**	model 3***	model 1*	model 2**	model 3***
EG(-1)	0.17132 (0.053)	----	----	0.124848 (0.0241)	----	----
industry(-1)	----	0.1396 (0.028)	----	----	0.616844 (0.0000)	----
manufacturing(-1)	----	----	0.05651 (0.154)	----	----	0.173883 (0.0000)
Provincial credits	2.2E (0.001)	0.06509 (0.0655)	0.00028 (0.0978)	9.56E (0.0821)	0.058326 (0.0000)	2.41E (0.0000)
Wage	0.0008 (0.010)	407566 (0.0000)	15004 (0.083)	5.75E (0.0007)	73.56827 (0.4961)	0.311513 (0.0000)
CPI	-0.00299 (0.0000)	-165692 (0.0177)	-179108 (0.0000)	-0.000825 (0.0000)	-126376.5 (0.2106)	74.55068 (0.0012)
Transport value added	2.87E-09 (0.0733)	0.9590 (0.0975)	0.9895 (0.0001)	1.73E-09 (0.0733)	0.232396 (0.0050)	0.000126 (0.0000)
Human Capital	2.27E (0.045)	1849.104 (0.0022)	116.24 (0.0288)	2.42E (0.0299)	1683.699 (0.0000)	0.241977 (0.3241)
Urban development	0.151 (0.540)	9.26E (0.0000)	6.16E (0.346)	1.43 (0.0000)	8.83E (0.0242)	296253.2 (0.0008)
GDP Percapita	0.0059 (0.181)	160922 (0.0142)	17667 (0.0000)	-3.93E (0.767)	115.81 (0.0022)	0.046886 (0.0001)
Government budget	1.42E (0.598)	0.2993 (0.840)	-0.05871 (0.930)	-3.93E (0.767)	4.1238 (0.0001)	0.000307 (0.0118)
J-Statistic	----	----	----	14.51065 (0.2692)	6.643467 (0.466921)	7.845729 (0.448683)
ρ	0.2008 (0.001)	-0.22071 (0.1814)	0.23781 (0.0541)	----	----	----
σ^2	0.001 (0.000)	4.08 (0.001)	5.55 (0.0001)	----	----	----

Note: () shows p-values.

* Dependent variable EG index

** Dependent variable industry value added

*** Dependent variable manufacturing value added

The results of spatial panel data and dynamic panel data are presented in Table (2). According to results the ρ estimates is 0.2, -0.22 and 0.23 for the model 1 to 3 respectively which is statistically significant for model 1 and 3 show the importance of a spatial

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autoregressive disturbance in the model and confirms the impact of cross-region spillovers on the spatial distribution of industrial activity. Lag dependent variable has included in the model to assess the impact of history of the region. According to the result in both estimations and models previous situation has a direct relationship with the current situation.

Generally, the signs of all the coefficients of the explanatory variables seem to be consistent with the theoretical expectations. Thus, the impact of provincial credit which measured by loan paid in each region on agglomeration forces seems to be positive throughout, so that it might be inferred that availability of credit have a significant potential for concentration of industrial activities. Hence, it might be asserted that degree of concentration might be increasing due to uneven distribution of banking paid loan in various regions. This results also in the GMM estimation are same. The positive relationship between the provincial credits and agglomeration can be explained by several reasons which have presented in details in the literature review.

Regional manufacturing wage rate and consumer price index as the welfare index of labor force and consumer are the determinant factor in the migration equation. As the results of spatial estimations in table (2) shows that higher regional manufacturing wage rate consistent with higher degree of concentration, while higher regional consumer price strengthen the dispersion forces. GMM estimation also confirm the results of spatial estimation with this difference that regional manufacturing wage rate in model 2 is not statistically significant and regional consumer price index in model 3 has a positive sign.

Trade freeness is an important factor in the firm location decision. For low trade freeness or high trade cost firms tend to a dispersed geographical distribution and for high trade freeness or low trade cost agglomeration is the long-run equilibrium. Regional share of transportation, post and telecommunication in GDP are used as a measure of trade freeness. Higher share consistent with the lower trade cost and thus higher agglomeration forces. The results in both spatial estimation and GMM estimation confirms a positive relationship between regional transportation value added and agglomeration indices.

Human capital, urban development, GDP Per capita has a positive effect on the concentration of industrial economic activity. While this variable can be seen to represent the development level of various province, it can be attractive to the workers and interpreted as the potential of the consumption for the industries. More importantly higher GDP per capita would have a chance of good access to the market or what called as home market effect (Combes and Overman, 2003). In the table (2) spatial estimation all this mentioned variables have the positive impact on the agglomeration but in the model 1 and 3 urban development and model 1 GDP per capita are not statistically significant. In the GMM estimation despite human capital in model 3 which is not statistically significant, all variables have the positive impact and consistent sign with the theoretical background.

5. Summary and Conclusion

In this paper, we examined the effect of monetary policy on the distribution of economic activity. The way which monetary policy affect the location decision is due to supply side and demand side. On the supply side credit availability in each region affect the decision of firm where to expand their production. On the demand side monetary policy affect the wage rate and price index which will have a great effect on the consumer welfare. We have developed a framework that combines labor migration and credit constraints and cost of living for consumers to study the role of monetary policy on the distribution of economic activity across regions. Results show the great impact of monetary policy on the distribution of economic activity where affect the credit availability of firms and welfare of consumer in various regions. In additions uneven distribution of credit across regions faced firms in less developed areas with critical performance challenges as to obtain external funds. This might lead them to limited source and higher cost of borrowing, higher risk of asset substitutions and inflexibility of capital structure which resulted in poorer performance and will change to a significant obstacle to their expansion, join new markets and export orientation production. It is important to say that the results of this survey do not imply that firms in less developed areas are sub-optimally located far from financial centers and metropolitan areas and should relocate their office to reduce debt costs

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and have better access to financial resources. In fact Policy makers must consider the fact that the uneven distribution of financial resources can have a serious impact on the economic performance of marginalized areas. Also it is important to say that financial issues are just one of many factors a firm has to take into account when selecting their location. This firms might favor their location because of specific industry geographical clustering, convenient access to intermediate inputs, proximity to suppliers or customers, state tax policy and considerations, or local favoritism by public officials, although access to financial resources can affect them substantially.

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